Statistics Review

FBE 464 1/10 &1/17 2001 Professor Hans O. Mikkelsen

Economic Theory

- Typically, you are given a theory like: $Y_t = X_t$
 - where \boldsymbol{Y}_t and \boldsymbol{X}_t are variables measured at time t.
- You want to know if empirical evidence:
 - Supports the theory.
 - With statistical confidence.
- Time series statistics!

Coefficients

• You want to test the theory, so you add a parameter and an intercept:

 $\mathbf{Y}_{t} = \mathbf{b}_{0} + \mathbf{b}_{1}\mathbf{X}_{t}$

• Economic theory specifies:

 $b_0 = 0$ and $b_1 = 1$

• Given some data you have to figure out what values of b₀ and b₁ are supported empirically.

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One observation

- Suppose in Jan. 2001 $Y_t=100$ and $X_t=100$.
- This implies that $Y_t = 0 + 1*X_t$ => Perfect fit!
- However, $Y_t = 100 + 0^*X_t$ is also a perfect fit.
- Therefore, we can not identify both b₀ and b₁ with only one observation.

Two observations

- Suppose in Dec. 2000 $Y_t=50$ and $X_t=50$, in Jan. 2001 $Y_t=100$ and $X_t=100$.
- Here there is only one solution: $Y_t = 0 + 1*X_t$
- The system is exactly identified and b₀=0 and b₁=1 is in accordance with theory.
- In general you need at least as many observations as you have parameters.

Many observations

- It is very imprecise to base anything on only two observations.
- Anything unusual happening during those obs., like an Asian crisis, is very unlikely to be informative about the whole population.
- Also, it can be shown that precision increases with the number of observations.
- Collect as much data as possible!

Independent observations

- With more observations than parameters you can not get a perfect fit, however.
 - Unless the observations are not independent (the previously mentioned two obs. are not because the second is just twice the first).
- Therefore, you expect to make an error at any point in time no matter how you choose the parameters.

Error term

- There are other reasons why you don't expect theory to hold at any point in time.
 - $-% \left(T^{\prime}\right) =\left(T^{\prime}\right) \left(T^$
 - Differences in timing.
 - Model is too simple (leaves something out).
- => Add an error term, ut:

$$\mathbf{Y}_{t} = \mathbf{b}_{0} + \mathbf{b}_{1}\mathbf{X}_{t} + \mathbf{u}_{t}$$

Estimation

- How can plausible values for the parameters be obtained, that is, how can the model be estimated?
- Sample of T observations on both variables. - $Y_1, Y_2, ..., Y_T; X_1, X_2, ..., X_T$
- It is easy to obtain parameters so that the average error is zero.

Estimating the model

- It turns out that the best (easiest) estimation procedure is OLS.
- Ordinary Least Squares (OLS).
 Minimizes the sum of squared residuals (errors)
 Minimizes the variance of regression errors.
- This avoids making too large negative or positive errors.

Estimation (cont.)

- Fortunately, if you include a constant the error term will have a mean of zero by construction.
- That is, not only do we not make too large errors, on average we make no error at all.
- There are other methods for estimation.
 All you need to know is that for your purpose OLS is best.

Example

• Excel can do OLS!

• Suppose the regression output is:

 $Y_t = 0.2 + 0.7*X_t + u_t$

- Do you conclude that theory fails because b₀ is not 0 and b₁ not 1?
- You don't know, there is uncertainty because we only observe a snapshot, say five years, and with noise.

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Standard errors

- Fortunately OLS (and Excel) also provides us with standard errors of the estimated parameters.
- The larger standard errors, the more uncertainty.
- If the standard error on b₁ is 0, there is no uncertainty and we can be 100% certain that b₁=0.7, and theory is rejected

Confidence intervals

- Assume we have many observations.
- Then, with 95% confidence we can say that a parameter is within an interval of plus, minus 2 std. dev. from its estimate.

• For
$$b_1$$
: 0.7 - 2*0.1 < b_1 < 0.7 + 2*0.1
=> 0.5 < b_1 < 0.9

• Because this interval does not include b₁=1, it is very likely that theory fails.

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Confidence intervals

- Assume that the standard error for the estimate of b_0 is 0.5.
- Then the 95% confidence interval for b_0 is: 0.2 - 2*0.5 $< b_0 < 0.2 + 2*0.5$

 $= -0.8 < b_0 < 1.2$

• Because this interval includes 0 it is very likely that b₀ is in accordance with theory.

T-test.

- It is more convenient to test theory using ttests.
- If you want to test if an estimated parameter is significantly different from a given value, compute:

 $t = (b-b^*)/s_b$

• where b is your estimate, b* is the value of interest, and s_b is the standard error of b.

t-test: decision rule

- You will find all you need from the regression output.
- Two sided test:
 - Reject the null hypothesis of $b=b^*$ if |t| is large.
- How large:
 - The t-test is t-distributed, and you can find a table in your favorite text book (or the handout).

t-test: example.

- Estimated b₁ is 0.7
- Estimated standard error of b₁ is 0.1
- You want to test $H_0: b_1=1.0$
- Test statistic is:
 - t = (0.7-1.0)/0.1 = -3.0
- If you have many obs., say 100, the 5% critical value is very close to 2.00

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t-test: example. (cont.)

- because |t|=3.00 > 2.00, you reject the null.
- The estimated parameter is not 1.
- Had the t-test been 1.5, you would not reject that b₁=1, since |t|=1.50 < 2.00.

t-test: How to find critical value

- A 5% significance level in a 2-sided test means that we have to cut 2.5% in both tails of the distribution.
- We find the 97.5% percentile to be approx. 2.0
- We know that a t-distribution is symmetric. so the 2.5% percentile is -2.0
- That's why |t| is a 5% test.

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Another example

- You want to test if the constant is significantly different from zero.
 - Estimated b_0 is 0.2
 - Std. error of \boldsymbol{b}_0 is 0.5
- t = (0.2-0.0)/0.5 = 0.4
- |t|=0.4 < 2.0 => Not statistically significant.
- Cannot reject H₀: b₀=0

Comparison

- Confidence intervals and t-tests tell exactly the same story.
- In fact, the reason why 2 is used as critical value for the t-test in large samples is the same as why we subtract and add 2 from the estimate to get a confidence interval.

Sub-set of parameters

- Suppose you have several explanatory variables in your model.
- Instead of testing hypotheses on some parameters individually, you may want to test a joint hypothesis like:
 - $-H_0: b_1 = b_2 = 0$
- Individual and joint test results may differ.
- Use an F-test.

Insignificant variables

- Get rid of them!
- In a finite sample, parameters that are not significantly different from zero:
 - Don't don't belong in the model.
 - Affects estimates of all other parameters.

=> <u>Only trust your results if all parameters</u> are significantly different from zero.

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\mathbb{R}^2

- R² is a measure of fit for the estimated model.
- R² = fraction of variation in Y_t that is explained by right hand side (RHS) variables.
- R²=0 means no relation at all between Y_t and RHS => Bad model.
- R²=1 indicates perfect fit => Perfect model.

R² - Problem

- PROBLEM: You don't know what a good fit is, and there is no test.
- R²=0.05 may be a reasonable fit for many financial models.
- R²=0.9 may not be good fit in other models

Plots

- To judge the fit of a model you can plot Y_t against the values predicted by the model.
- Y_t will always be more volatile than predictions.
- However, maybe the predictions capture the direction of most changes in Y_t.
- Always trust tests before R² and plots.

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Bad news

- The bad news is that OLS only works optimally if the error term (residual) is so-called "white noise".
- There are three conditions for white noise: 1) Mean zero: $E_t[u_{t+1}] = 0$, all t

2) Constant variance: $E_t[u_{t+1}^2] = s^2$, all t

3) No serial correlation: $E_t[u_{t+k}|u_{t-n}] = 0$, all t,n,k

Conditions

- The 1st condition is satisfied when using OLS, and including a constant.
- Violations of the 2nd condition are very difficult to handle.
 - => Ignore them for the purpose of this course!

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Serial correlation

- Violations of the 3rd condition are <u>very</u> serious!
- However, you can typically detect and correct for serial correlation.
- Often you find first order serial correlation:
 E_t[u_{t+1}|u_t] is different from zero.

Ex.: 1st order serial correlation

• $E_t[u_{t+1}|u_t] = 0.5*u_t$

=> The correlation between u_t and u_{t+1} is 0.5

- Alternatively, the first order autocorrelation coefficient of the residuals is 0.5.
- E.g., if the current error is 10% we expect the next error to be 0.5*10%=5%.

Implications of ser. corr.

- If the models prediction was 10% above the actual outcome, we expect it to be 5% above next period.
- That is, given this periods error we can forecast next period's error.
- The model fails to take some structure into account when there is serial correlation.
 Model could and should be improved.

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Rational expectations

- Often u_t is a rational expectations error. – Therefore, there cannot be serial correlation.
- If there was, I would be making systematic expectations errors. NOT rational!
 - If I'm 10% wrong today it is not rational to know that I will be 5% wrong tomorrow, without taken the information into account!

Statistical implications

• If there is serial correlation, not just first order, standard errors of parameter estimates are wrong.

=> You can't make correct inference.

All we do involves testing some hypotheses.
 Without correct inference this is an impossible task.

Durbin-Watson test

- Test for <u>first</u> order serial correlation.
- Null hypothesis: no serial correlation.
- DW = 2(1-r₁) where r₁ is the first order autocorrelation coefficient of the residuals.
- The above expression is an approximation.
- Critical values for DW can be found in almost any introductory time series book.

Durbin-Watson: Rules

- High first order serial correlation means: DW=2(1-r₁) is small.
- Hence, small values of DW rejects the null hypothesis of no serial correlation.
- If DW is sufficiently low you have serial correlation.
- If DW is sufficiently high you don't.

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Durbin-Watson: example

- First order autocorrelation is 0.5 => DW = 2(1-0.5) = 1.0
- According to my textbook:
 - DW_L=1.65 and DW_U=1.69 for our sample size of T=100, and 2 right hand side variables in my regression.

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• Since DW is lower than the lowest critical value, we have detected serial correlation.

Other examples

- Had DW been between the two critical values, say 1.67, the test would be inconclusive.
- Had DW been above the upper critical value, say 2.00, we would have no serial correlation.

Notice:

- The described DW procedure is a test against positive serial correlation.
 - r₁ is positive.
- If you want to test negative serial correlation you should use: 4-DW.
- I.e. if 4-DW is sufficiently low, you have a problem of negative first order serial correlation.

Correcting for serial correlation

- If you detect serial correlation in the residuals, theory may already be wrong! E.g., rational expectations ass. may be wrong.
- Otherwise, re-specify your model.
 - Include a lagged dependent variable on the RHS:
 - $\mathbf{Y}_{t} = \mathbf{b}_{0} + \mathbf{b}_{1}\mathbf{X}_{t} + \mathbf{b}_{2}\mathbf{Y}_{t\text{-}1} + \mathbf{u}_{t}.$
- Often the serial correlation goes away.

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Serial correlation persists

- Re-specify again!
 - $-\ensuremath{\text{ include more lags of }} Y$ on the right hand side.
- If this doesn't work, transform your variables:
 - For example, estimate the model in differences.
- If nothing works there are only complicated methods:
 - Not for FBE 464!

Trending Variables.

- Variables that "explodes" over time.
- Maybe they:
 - get very, very high values.
 - just continue to increase.
 - $-\ensuremath{\, \text{never}}$ seem to revert towards a natural level.
- That gives us special statistical and numerical problems. <u>Variables are non-</u><u>stationary and need to be transformed.</u>

Log of variables

- Instead of the variable itself, e.g. Y, we often work on ln(Y).
 - => Very large numbers get smaller.
- Also, macro and financial variables are typically log normally distributed (roughly).
 => ln(variable) is normally distributed

Growth rates

- One way to get rid of the trends discussed before is by working on growth rates: Growth rate 1 of $Y_t = (Y_t - Y_{t-1})/Y_{t-1}$
- Since we often use ln(Y_t), a more convenient way to compute growth rates is: Growth rate 2 of Y_t = ln(Y_t)-ln(Y_t)

2 types of growth rates

• If growth rates are small, growth rates 1 and 2 are very similar.

• If
$$Y_{t-1}=55$$
 and $Y_t=56$

However, if
$$Y_{t-1}=55$$
 and $Y_t=100$

$$=> G1 = (100-55)/55 = 0.8182$$

 $G2 = \ln(100) - \ln(55) = 0.5978$

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Use log growth rates

- Suppose Y_{t-2} =100, Y_{t-1} =200, and Y_t =100.
- Then G1_{t-1}=1.0 and G1_t=-0.5.
 => You will conclude that Y changes by an average of 0.25 or 25% per period, although it ends up unchanged!
- G2_{t-1}=0.69 and G2_t=-0.69.
 => You conclude that on average Y does not change at all, which is the truth!

The End.	
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